

Lecture 23

Clustering (4)

Outline

- Clustering as Density Estimation
- Mixture Density Estimate
- Non-parametric density estimate: Parzen Windows

Clustering and Density Estimate

- Probability density function $p(x)$ describes the distribution of samples in the feature space.
- The clustering results $w_m(t) \propto p(x)$
- $p(x)$ can also be estimated using
 - Parametric method: mixture density model
 - Non-parametric method: parzen window

Mixture Density Estimation

$$p(\mathbf{x}|\theta) = \sum_{i=1}^c p(\mathbf{x} | \theta(i), i) p(i)$$

- Assume # of clusters "c" is known. Want to estimate the Prior Probabilities $\{p(i); 1 \leq i \leq c\}$, and $\theta = [\theta(1), \dots, \theta(c)]$.
- Let $\{x_k; 1 \leq k \leq n\}$ be drawn from the mixture density, Log-Likelihood $L(\mathbf{x}) = \log \left\{ \prod_{k=1}^n p(x_k | q) \right\} = \sum_{k=1}^n \log p(x_k | \theta)$

$$\begin{aligned} \nabla_{\theta(i)} L(\mathbf{x}) &= \sum_{k=1}^n \frac{1}{p(x_k | q)} \{ \nabla_{\theta(i)} p(x_k | \theta(i), i) p(i) \} \\ &= \sum_{k=1}^n \frac{p(x_k | q(i), i) p(i)}{p(x_k | q)} \nabla_{\theta(i)} [\log p(x_k | \theta(i), i)] \end{aligned}$$

Mixture of Gaussian Densities

- $p(\mathbf{x}|\mu(i), \Sigma(i), i) = (2\pi|\Sigma|)^{1/2}\exp\{-[|\mathbf{x}-\mu(i)|^T\Sigma^{-1}|\mathbf{x}-\mu(i)|]/2\}$
- θ (μ and Σ) can be estimated as:

$$\hat{\mu}(i) = \frac{\hat{p}(i | x_k, \hat{\mu}(i), \hat{\Sigma}(i))}{\sum_{k=1}^N \hat{p}(i | x_k, \hat{\mu}(i), \hat{\Sigma}(i))} \mathbf{x}_k$$

$$\hat{\Sigma}(i) = \frac{\hat{p}(i | x_k, \hat{\mu}(i), \hat{\Sigma}(i))}{\sum_{i=1}^N \hat{p}(i | x_k, \hat{\mu}(i), \hat{\Sigma}(i))} (\mathbf{x}_k - \hat{\mu}(i))(\mathbf{x}_k - \hat{\mu}(i))^T$$

- Parameters appear in both sides of the equations \Rightarrow iterative solution is required.

ML Estimates of Mixture Parameters

- ML estimate of class i prior probability:

$$\hat{p}(i) = \frac{1}{n} \sum \hat{p}(i | \mathbf{x}_k, \hat{\mu}(i), \hat{\Sigma}(i));$$

$$\begin{aligned} \hat{p}(i | \mathbf{x}_k, \hat{\mu}(i), \hat{\Sigma}(i)) &= \frac{p(\mathbf{x}_k | i, \hat{\mu}(i), \hat{\Sigma}(i)) \hat{p}(i)}{\sum_{j=1}^c p(\mathbf{x}_k | j, \hat{\mu}(j), \hat{\Sigma}(j)) \hat{p}(j)} \\ &= \frac{|\hat{\Sigma}(i)|^{-1/2} \exp\{-\frac{1}{2} [(\mathbf{x}_k - \hat{\mu}(i))^T \hat{\Sigma}^{-1}(i) (\mathbf{x}_k - \hat{\mu}(i))]\} \hat{p}(i)}{\sum_{j=1}^c |\hat{\Sigma}(j)|^{-1/2} \exp\{-\frac{1}{2} [(\mathbf{x}_k - \hat{\mu}(j))^T \hat{\Sigma}^{-1}(j) (\mathbf{x}_k - \hat{\mu}(j))]\} \hat{p}(j)} \end{aligned}$$

Parzen Windows

- A non-parametric method using interpolation functions
 - Let $\varphi(u)$ be such that $\varphi(u) \neq 0$ and $\int \varphi(u) du = 1$
- $\{\mathbf{x}(k); 1 \leq k \leq N\}$ are drawn from unknown density $p(\mathbf{x})$.
Set $V_N = h_N^d$ where h_N is a smoothing parameter and d is the dimension of \mathbf{x} . Then the estimate of $p(\mathbf{x})$ is:

$$\hat{p}_N(\mathbf{x}) = \frac{1}{NV_N} \sum_{k=1}^N \varphi(\mathbf{x} - \mathbf{x}(k)) / h_N$$

- Conditions: $\lim_{n \rightarrow \infty} V_N = 0$, and $\lim_{n \rightarrow \infty} NV_N = \infty$.
- Example of window function: $\varphi(u) = \frac{1}{\sqrt{2\pi}} \exp(-u^2/2)$